

1 **Re:** Effect of hydroxychloroquine with or without azithromycin on the mortality of COVID-  
2 19 patients: a systematic review and meta-analysis

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11 To the Editor,

12 We read with interest the meta-analysis published in *Clinical Microbiology and Infection* by  
13 Fiolet *et al.* [1] entitled "Effect of hydroxychloroquine with or without azithromycin on the  
14 mortality of COVID-19 patients: a systematic review and meta-analysis". This meta-analysis  
15 concluded that the combination of hydroxychloroquine (HCQ) and azithromycin (AZ) was  
16 associated with increased mortality and that HCQ alone had no effect on mortality.

17 We believe that this study is fatally flawed. To start with, its conclusion is impossible in light  
18 of the treatment in our center of patients with HCQ-AZ, with reduction of mortality in the  
19 population at risk (>60 years) by a factor two [2]. In this context, we sought to understand  
20 how the authors reached their conclusions.

21 First, none of the authors has extensive experience in the treatment of infectious diseases, and  
22 a generic systematic review of literature does not replace expert understandings of study  
23 methods and pitfalls, as we have described [3]. The authors report several meta-analyses but  
24 omitted ours [4]. Our analysis highlighted a major factor of heterogeneity between studies:  
25 Rote "Big-data" studies, with complete disconnection between data analysis and clinician  
26 expertise, and clinical studies, in which the analysis is conducted by clinicians with a precise  
27 and detailed protocol. Our work notably contributed to highlighting the Lancetgate [5] and  
28 showed that Big-data studies are associated with conflicts of interest and absence of detailed  
29 treatment protocols.

30 The fatal flaw of the Fiolet *et al.* analysis is that it used subjective, capricious and specious  
31 criteria decisions about which studies to include. Large valid observational studies reporting  
32 significant benefit and published during the inclusion period and that used standard accepted  
33 methods to control for confounding factors (propensity-score matching) were not included,  
34 notably Arshad *et al.* in the USA (n = 2,541), Bernaola *et al.* in Spain (n = 1,645), and our  
35 study of 199 patient pairs in France.

36 One inclusion criterion mentioned by the authors [1] is “cases confirmed by RT-PCR”. This  
37 however is in contradiction to inclusion of Skipper *et al.*, with “Only 58% of participants  
38 received SARS-CoV-2 testing” and the RECOVERY trial for which PCR confirmation was  
39 not mandatory, as well as that it used toxic doses (2400 mg HCQ within the first 24 hours).  
40 Worse, Fiolet *et al.* included data from the study by Rivera *et al.* which itself is fatally flawed.  
41 That study, of cancer patients, included “Participation by anonymous individual health-care  
42 practitioners located in Argentina, Canada, the EU, the UK, and the USA is also allowed. The  
43 mechanism of data collection can be retrospective (after the course of COVID-19) or  
44 concurrent, at the discretion of the respondent.” This is not a sampling frame for any type of  
45 epidemiologic study. There is no assurance that individual practitioners didn't select patients  
46 on which to report because the patients didn't do well on antiviral therapies. Haphazard  
47 subject collection is not a representative epidemiologic sampling method and has no  
48 justification for use. Second, the Rivera *et al.* data show dramatic differences in HCQ and AZ  
49 use for nontreated vs treated subjects by baseline disease severity, and the authors did not  
50 report results on HCQ+AZ use but on HCQ+other medication use, which is not adjusted  
51 adequately for severity. Simply put, patients with worse conditions were given more  
52 medications and were more likely to die of their cancers. Third, Fiolet *et al.* use results values  
53 in their forest plot that do not appear in the paper or supplement of Rivera *et al.*  
54 Two new retrospective studies further contradict the authors' conclusions. The study of 8,075  
55 patients in Belgium (Catteau, 2020) and 3,451 patients in Italy (Di Castelnuovo, 2020) report  
56 clear benefits of chloroquine derivatives on mortality (Figure). In this context, we present  
57 (Figure) an update of our meta-analysis [4]. The clinical studies include 4,121 patients from 7  
58 countries, with summary odds ratio 0.60, 95% confidence interval 0.50-0.73,  $p=10^{-7}$ ). Results  
59 were consistent among studies ( $I^2=20\%$ ,  $p=0.23$ ). Among the 35,985 patients included in the

60 Big data studies, a significant but smaller benefit was found (0.84, 0.75-0.94,  $p=.003$ ) with  
61 appreciable heterogeneity ( $I^2=76\%$ ,  $p<.001$ ).

62 History and reasoning demonstrates that the authors' conclusions are erroneous. We suggest  
63 that the authors leave their political motivations at the door before undertaking scientific  
64 work.

65

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70 **Declaration of competing interest**

71 The authors declare no competing interests. Funding sources had no role in the design and  
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73 preparation, review, or approval of the manuscript. Our Marseille group used widely available  
74 generic drugs distributed by many pharmaceutical companies.

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95 **Figure legend**

96 **Figure. Meta-analysis on hydroxychloroquine for COVID-19 mortality (August 26,**  
97 **2020)**

98 A. Studies published online during the inclusion period of Fiolet *et al.* [1] (before July, 25)  
99 but not included in their work are in red. Arshad, Lagier and Bernaola used propensity score  
100 matching (PSM). \*Studies published after July 25, 2020. CI: confidence interval, HCQ:  
101 hydroxychloroquine, RCT: randomized controlled trial. This meta-analysis was performed  
102 with a random-effects model using Comprehensive Meta-Analysis v3 (Biostat, Englewood,  
103 NJ, USA). B. Studies included in Fiolet *et al.* [1] but excluded in the present analysis and  
104 reasons for exclusion. aOR: adjusted Odds ratio.